

# Mining Data Streams-Estimating Frequency Moment

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- ▶ Let  $f_i$  be the number of occurrences of the  $i$ th element for any  $i \in [1, n]$ , then the  $k$ th frequency moment is  $F_k = \sum_i f_i^k$

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## Frequency Moment

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- ▶ The 1st moment is the sum of the  $f_i$ 's which must be the length of the stream. This is easy to calculate.
- ▶ The 2nd moment is the sum of the squares of the  $f_i$ 's. It is sometimes called the *surprise number* as it measures the unevenness of the distribution of elements.
  - ▶ Suppose we have a stream of length 100.
  - ▶ Scenario 1: There are 10 elements each with frequency 10.  
$$F_2 = 10 * 10^2 = 1000$$
  - ▶ Scenario 2: There are 10 elements, 1st item has frequency 91, and rest have each frequency 1.  $F_2 = 91^2 + 9 * 1^2 = 8290$ .

## Computing $F_2$ in Small Space

- ▶ Linear Sketching
- ▶ Alon-Matias-Szegedy Sampling (read Sec 4.5 Leskovec et al.)

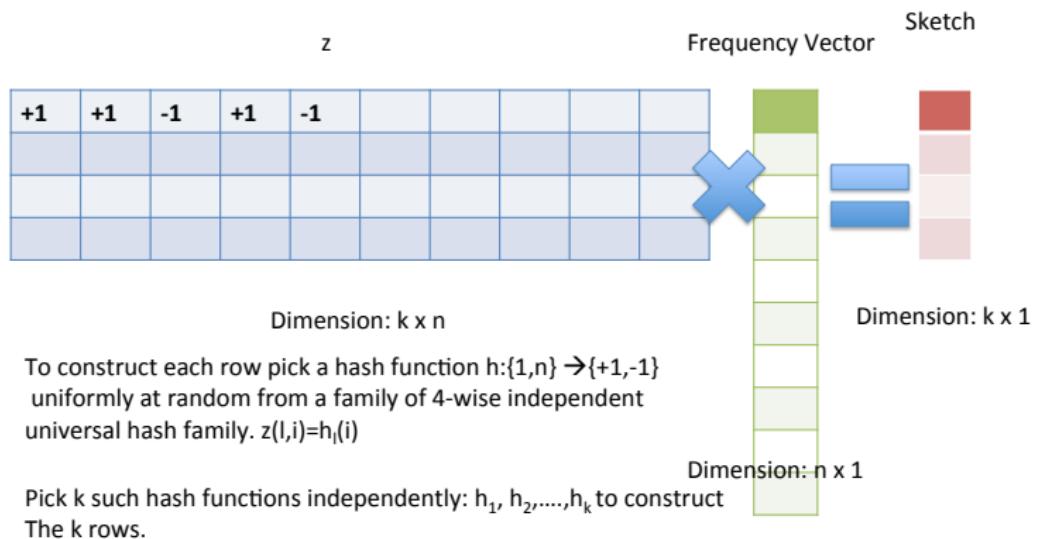
## Linear Sketch for $F_2$

- ▶ **Problem** Given a stream  $A_1, A_2, \dots, A_m$  where elements are coming from the universe  $[1, n]$  estimate  $F_2 = \sum_{i=1}^n f_i^2$  in “small space”.
- ▶ **Output** Return an estimate  $\hat{F}_2$  such that

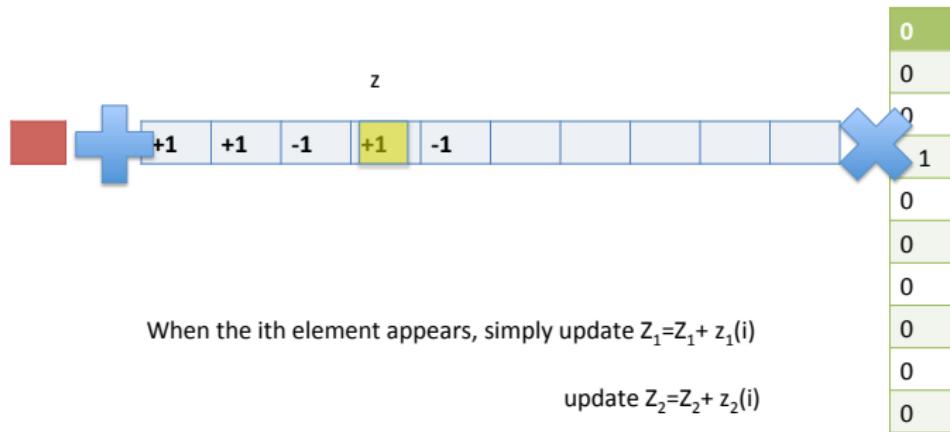
$$\Pr \left( F_2(1 - \epsilon) \leq \hat{F}_2 \leq (1 + \epsilon)F_2 \right) \geq (1 - \delta)$$

where  $\epsilon > 0$  and  $\delta > 0$  are respectively the error and confidence parameters.

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When the  $i$ th element appears, simply update  $Z_1 = Z_1 + z_1(i)$

update  $Z_2 = Z_2 + z_2(i)$

update  $Z_3 = Z_3 + z_3(i)$

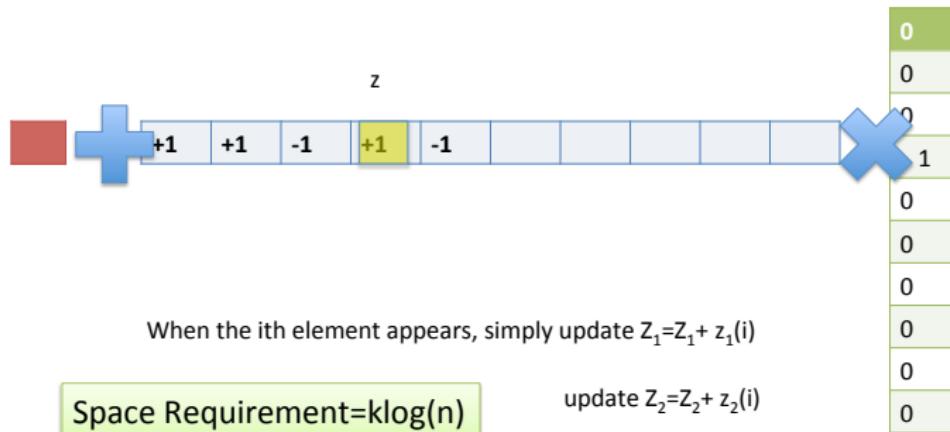
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$$\text{Estimate} = (Z_1^2 + Z_2^2 + \dots + Z_k^2) / k$$

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- ▶ Apply Chebyshev.

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# Expectation of $Z_s^2$

$Z_s \sim Z, s = 1, 2, \dots, k$

- ▶  $Z = \sum_{i=1}^n f_i z(i), Z^2 = \sum_{i,j \in [1,n]} f_i f_j z_i z_j$

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since  $E[z_i z_j] = 0$  if  $i \neq j$  and  $E[z_i^2] = 1$ .

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$$E[\hat{F}_2] = \frac{1}{k} \sum_{s=1}^k E[Z_s^2] = F_2$$

## Variance of $Z_s^2$

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since  $E[z_i z_j z_k z_l] = 0$  if  $i < j < k < l$  or 3 of the terms are equal.

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$$Var(Z^2) = 4 \sum_{i,j:i < j} f_i^2 f_j^2 \leq 2 F_2^2$$

## Variance of $\hat{F}_2$

$$\text{Var}(\hat{F}_2) = \text{Var}\left(\frac{1}{k} \sum_{s=1}^k Z_s^2\right)$$

$$= \frac{1}{k^2} \text{Var}\left(\sum_{s=1}^k Z_s^2\right) \text{ since } \text{Var}(aX) = a^2 \text{Var}(X) \text{ for any constant } a$$

$$= \frac{1}{k^2} \sum_{s=1}^k \text{Var}(Z_s^2) \leq \frac{1}{k^2} 2kF_2^2 = \frac{2F_2^2}{k}$$

## Boosting Confidence by Median

- ▶ We have

$$\text{Prob} \left( F_2(1 - \epsilon) \leq \hat{F}_2 \leq (1 + \epsilon)F_2 \right) \geq \frac{7}{8}$$

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- ▶ Take  $t$  independent estimates

$$H_1 = \hat{F}_2^1, H_2 = \hat{F}_2^2, \dots, H_t = \hat{F}_2^t$$

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- ▶ Return the median of  $H_1, H_2, \dots, H_t$ .

## Boosting by Median

- ▶ Suppose there is an Algorithm that returns an estimate  $\hat{F}$  of a true estimate  $F$  such that  $|\hat{F} - F|$  is small with probability  $\frac{7}{8}$ .
- ▶ How can we design an algorithm that will return an estimate  $G$  of  $F$  such that  $|G - F|$  is small with probability  $99/100$ ?  
(In general  $1 - \delta$ )

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- ▶ How can we design an algorithm that will return an estimate  $G$  of  $F$  such that  $|G - F|$  is small with probability  $99/100$ ? (In general  $1 - \delta$ )
- ▶ Run  $s = 6 \log \frac{1}{\delta}$  independent copies of the Algorithm to obtain estimates  $\hat{F}^1, \hat{F}^2, \dots, \hat{F}^s$ . Set  $G = \text{median}_{i=1}^s \hat{F}^i$ .

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- ▶ Bound

$$\text{Prob}(Y > 3 \log \frac{1}{\delta})$$

using Chernoff's bound.

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- ▶ Upper Tail version of Chernoff Bound. For  $\epsilon > 1$

$$\text{Prob}(Y > E[Y](1 + \epsilon)) \leq e^{-\frac{E[Y]\epsilon^2}{2+\epsilon}}$$

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$$\begin{aligned} \text{Prob}\left(Y > 3 \log \frac{1}{\delta}\right) &= \text{Prob}\left(Y > \frac{3}{4} \log \frac{1}{\delta}(1 + 3)\right) \\ &\leq e^{-\frac{3}{4} \left(\log \frac{1}{\delta}\right) 9 \frac{1}{5}} < \delta \end{aligned}$$

# Versions of Chernoff Bound

Reference:

<https://www.cs.princeton.edu/courses/archive/fall09/cos521/Handouts/probabilityandcomputing.pdf>

# Frequency Moment

- ▶ For  $k > 2$ , the best bound known is  $\tilde{O}(n^{1-\frac{2}{k}} \log \frac{1}{\delta})$  barring  $\text{poly}(\frac{1}{\epsilon})$  factor. There is an almost matching lower bound of  $\Omega(n^{1-\frac{2}{k}})$ .
- ▶ For  $k < 2$ , the best bound known is  $\tilde{O}(\frac{1}{\epsilon^2} \log \frac{1}{\delta})$ .
- ▶ The algorithms use clever combination of sketching and hashing

# Sketching as a Versatile Tool

- ▶ Estimating entropy, quantiles, heavy hitters, fitting histograms etc.
- ▶ Applications beyond streaming: dimensionality reduction, nearest neighbors, anomaly detection, statistics over social network.
- ▶ Not only useful for small-space algorithm design, but also for fast running time, distributed processing etc.

## A different linear sketch

- Instead of  $\pm 1$ , let  $r_i$  be i.i.d. random variables from  $N(0,1)$
- Consider

$$Z = \sum_i r_i x_i$$

- We still have that  $E[Z^2] = \sum_i x_i^2 = \|x\|_2^2$ , since:
  - $E[r_i] E[r_j] = 0$
  - $E[r_i^2] = \text{variance of } r_i, \text{ i.e., } 1$
- As before we maintain  $Z = [Z_1 \dots Z_k]$  and define  
$$Y = \|Z\|_2^2 = \sum_j Z_j^2 \quad (\text{so that } E[Y] = k\|x\|_2^2)$$
- We show that there exists  $C > 0$  s.t. for small enough  $\epsilon > 0$

$$\Pr[ |Y - k\|x\|_2^2| > \epsilon k\|x\|_2^2] \leq \exp(-C \epsilon^2 k)$$

Slide from Piotr Indyk's course on Streaming, Sketching and Compressed Sensing

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